Animal Behaviour 151 (2019) 29-42



Contents lists available at ScienceDirect

Animal Behaviour

journal homepage: www.elsevier.com/locate/anbehav

North American red squirrels mitigate costs of territory defence through social plasticity



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ARTICLE INFO

Article history: Received 20 May 2018 Initial acceptance 16 July 2018 Final acceptance 11 January 2019

MS. number: A18-00328R3

Keywords: behavioural plasticity behavioural time budget dear enemy familiarity red squirrel social competence social environment territoriality vocalization For territorial species, the ability to be behaviourally plastic in response to changes in their social environment may be beneficial by allowing individuals to mitigate conflict with conspecifics and reduce the costs of territoriality. Here we investigated whether North American red squirrels, Tamiasciurus hudsonicus, are able to minimize costs of territory defence by adjusting behaviour in response to the familiarity of neighbouring conspecifics. Since red squirrels living in familiar neighbourhoods face reduced intrusion risk, we predicted that increasing familiarity among territorial neighbours would allow squirrels to spend less time on territorial defence and more time in the nest. Longitudinal behavioural data (1995-2004) collected from the same squirrels across several different social environments indicated that red squirrels reduced rates of territorial vocalizations and increased nest use in response to increasing familiarity with neighbours. In contrast, cross-sectional data (2015-2016), which provided observations from each individual in a single social environment, did not provide evidence of this plasticity. Post hoc analyses revealed that evidence of social plasticity in this system is primarily due to within-individual changes in behaviour, which we were unable to estimate in the cross-sectional data. Our results demonstrate that red squirrels respond to changes in their social environment by adjusting their behaviour in a manner that reduces the costs of territoriality. However, our results also suggest that estimating plasticity by comparing behaviour among individuals (i.e. cross-sectional analyses) may not always be reliable. Our ability to detect these effects may therefore depend on having data with multiple observations from the same individuals across different social environments.

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Phenotypic plasticity can broadly be defined as the ability of a single genotype to express multiple phenotypes in response to different environmental conditions (Pigliucci, 2001). Indeed, classic studies of phenotypic plasticity have focused on changes in nonreversible traits (e.g. morphology) that are expressed within a single genotype (Greene, 1989; Hebert & Grewe, 1985; Lively, 1986; but see Herzog, Tittgen, & Laforsch, 2016, for an example of when morphological traits can be reversible). However, traits that are expressed repeatedly over the course of an organism's lifetime (e.g.

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timing of reproduction) can be subject to reversible withinindividual plasticity (Nussey, Wilson, & Brommer, 2007; Piersma & Drent, 2003). This 'reversible plasticity' (Gabriel, Luttbeg, Sih, & Tollrian, 2005), also referred to as 'phenotypic flexibility' (Piersma & Drent, 2003), or 'responsiveness' (Wolf, Van Doorn, & Weissing, 2008), is a powerful mechanism for adapting to changing and unpredictable environmental conditions. Behavioural traits, in particular, show capacity for substantial phenotypic lability in response to changing environmental conditions within an organism's lifetime. This behavioural flexibility is one form of phenotypic plasticity that can facilitate an organism's ability to cope with both predictable and unpredictable variation in the environment (Ghalambor, Angeloni, & Carroll, 2010).

https://doi.org/10.1016/j.anbehav.2019.02.014

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The social realm is potentially one of the most dynamic and variable aspects of an individual's environment, since high levels of unpredictability are inherent when interacting with other agents that can also exhibit plasticity in behaviour. Examples of social plasticity (changes in behaviour in response to changing social conditions; Montiglio, Wey, Chang, Fogarty, & Sih, 2017; Sih, Chang, & Wey, 2014) are widespread. For instance, individuals adjust their level of aggression according to the perceived level of threat imposed by neighbours versus strangers (Temeles, 1994). Interacting individuals change their signalling behaviour in response to bystanders ('audience effect': Doutrelant, McGregor, & Oliveira, 2001; Pinto, Oates, Grutter, & Bshary, 2011). Behaviour may also be affected by previous social experiences such as 'winner-loser effects' (Hsu, Earley, & Wolf, 2006; Rutte, Taborsky, & Brinkhof, 2006), as well as by 'eavesdropping', in which bystanders extract information from interacting individuals (Earley, 2010; Mennill, Ratcliffe, & Boag, 2002; Oliveira, McGregor, & Latruffe, 1998).

The ability to adjust behaviour in response to social context should allow individuals to avoid costly interactions while appropriately engaging in other social interactions that might enhance fitness (Taborsky & Oliveira, 2012). This ability to show adaptive adjustments in social behaviour has been termed 'social skill' (Sih & Bell, 2008) or 'social competence' (Taborsky & Oliveira, 2012), although only a few studies have directly demonstrated fitness benefits of social plasticity (e.g. Han & Brooks, 2015; Montiglio et al., 2017; Patricelli, Uy, Walsh, & Borgia, 2002). Given the substantial number of social interactions that group-living species must navigate, the benefits of social plasticity are expected to be high in such species (Taborsky, Arnold, Junker, & Tschopp, 2012). However, solitary, territorial species may also benefit from appropriate adjustments in social behaviour, as being socially plastic may allow individuals to mitigate conflict with conspecifics and reduce the costs of territoriality. For example, gladiator treefrogs, Hypsiboas rosenbergi, adjust the timing of vocalizations in response to changing levels of conspecific competition. By reducing calling rates in response to changing social conditions, individuals can minimize an energetically costly behaviour (Höbel, 2015).

Solitary, territorial species, like their social counterparts, face variation in their social environments through their interactions with territorial neighbours. A well-described example of this variation is differences in familiarity with neighbours (Bebbington et al., 2017; Beletsky & Orians, 1989; Eason & Hannon, 1994; Grabowska-Zhang, Wilkin, & Sheldon, 2012). Long-term social relationships with neighbours have been presumed to be advantageous by minimizing renegotiation of territory boundaries and therefore reducing aggression as well as time and energy spent on territory defence ('dear enemy effect'; Fisher, 1954). However, most evidence in support of this phenomenon comes from experimental studies where individuals are exposed to a familiar and unfamiliar stimulus and a behavioural response is recorded (Temeles, 1994). We know less about how behavioural time budgets are affected by long-term social relationships under natural conditions when individuals may have to navigate territorial dynamics with multiple neighbours (but see Bebbington et al., 2017; Eason & Hannon, 1994).

In this study, we examined whether territorial North American red squirrels, *Tamiasciurus hudsonicus* (hereafter 'red squirrels'), adjust their behaviour in response to their familiarity with their local social environment. Red squirrels are territorial rodents that defend year-round exclusive territories (Smith, 1968). In the Yukon, red squirrels cache white spruce, *Picea glauca*, cones in a larder hoard called a 'midden' at the centre of their territory (Fletcher et al., 2010). This food cache is important for overwinter survival (Kemp & Keith, 1970; LaMontagne et al., 2013) and both sexes heavily defend these resources from conspecifics, primarily through territorial vocalizations called 'rattles' (Smith, 1978). Rattles function to deter intruders (Siracusa, Morandini, et al., 2017b) but are also individually unique (Digweed, Rendall, & Imbeau, 2012; Wilson et al., 2015). Rattles therefore carry important information about the local social environment, such as the identity or density of neighbouring conspecifics. Squirrels use this acoustic information to increase rattling rates and vigilance and decrease nest use in response to increasing local density (Dantzer, Boutin, Humphries, & McAdam, 2012), providing some evidence of functional plasticity in territorial behaviour. Additionally, there is evidence that local social conditions are temporally variable in this system. Overturn of middens can occur through the death of a territory owner or through bequeathal. As a result, some squirrels may occupy different territories each year, leading to variation in neighbour familiarity (i.e. duration of tenure as neighbours). This can affect local territory conditions. Familiarity with territorial neighbours has been shown to have direct effects on territory intrusion risk. Specifically, individuals living in neighbourhoods with higher average familiarity faced reduced intrusion risk (Siracusa, Boutin, et al., 2017a), consistent with the dear enemy phenomenon (Fisher, 1954).

Given temporal heterogeneity in territorial neighbours and variation in signalling behaviour, we predicted that increasing familiarity with territorial neighbours would allow for changes in other aspects of behaviour, specifically decreased time spent on territorial defence as evidenced by (1) decreasing rattling rates and (2) reducing time spent vigilant for conspecifics. We also predicted that, as a squirrel's familiarity with its neighbours increased, squirrels would increase the proportion of time spent in the nest, as a proxy for time spent on offspring care or self-maintenance. Changes in behaviour, as predicted above, would allow individuals to minimize aggression and reduce allocation of time and energy to territory defence under social conditions associated with reduced risk of territorial intrusion, and thus would be indicative of social competence in this species.

METHODS

We studied a natural population of North American red squirrels located in the southwest Yukon near Kluane National Park (61°N, 138°W) that has been monitored continuously since 1987 as part of the Kluane Red Squirrel Project (KRSP; McAdam, Boutin, Sykes, & Humphries, 2007). To assess social plasticity in red squirrels, we measured behaviour of individuals on three study grids characterized by open boreal forest where white spruce is the dominant tree species (Krebs, Boutin, & Boonstra, 2001).

In this study, we used a longitudinal data set spanning 8 years and cross-sectional data from 2 years to assess changes in behaviour. Our longitudinal data set contained multiple observations of the same individuals across different social environments, while our crosssectional data represented an intensive snapshot of a large number of individuals at a single point in time (i.e. a single social environment for each individual). Our cross-sectional data, therefore, only allowed us to infer plasticity from differences in behaviour among different individuals experiencing different environments. Although behavioural plasticity is fundamentally a within-individual phenomenon, it can be approximated by comparing among individuals in different environments (Legagneux & Ducatez, 2013; Slabbekoorn & Peet, 2003). While this among-individual approach is a useful tool (particularly where it is challenging or time consuming to collect data on many individuals over several environments), it relies on the critical assumption that the among-individual relationship is an accurate representation of within-individual changes in behaviour.

Our longitudinal data set included long-term focal animal observations (Altmann, 1974) of 41 red squirrels across 8 years (1995–2004), collected on one unmanipulated control grid (Sulphur: SU; 40 ha). On average, we had data for two social environments per individual (range 1–8 social environments), meaning that our longitudinal data contained multiple observations across different social environments for most, but not all, individuals. Analyses using the longitudinal data therefore incorporated variation that was due to within-individual changes in behaviour as well as variation due to differences in behaviour among individuals.

Our cross-sectional data included focal observations of 108 squirrels in 1 year (2016) on two unmanipulated control grids (Kloo: KL and SU; 40 ha each) and one food-supplemented grid (Agnes: AG; 45 ha; see Dantzer et al., 2012, for a description of the food supplementation experiment). In this cross-sectional data we only had observations from each individual in a single social environment and, therefore, could only estimate social plasticity by comparing changes in behaviour among individuals. Since accurately capturing behavioural differences is often challenging, even with the intensive use of focal observations, we also measured the behaviour of squirrels by deploying accelerometers in 2016 to assess nest use, and audio recorders in 2015 and 2016 to measure rattling rates. All audio and accelerometer data collected in 2016 were from a subset of the same 108 individuals that we conducted focal observations on that year. All focal observations, audio data and accelerometer data were collected between May and September, since this is the time during which we regularly monitor the red squirrel population and have detailed information on territory ownership. Further details on these approaches are provided below.

Measuring Familiarity

In each year, we enumerated all squirrels living on our study areas and monitored individuals from March until August. We used a combination of live-trapping procedures and behavioural observations to track reproduction, identify territory ownership, and determine offspring recruitment from the previous year (see Berteaux & Boutin, 2000; McAdam et al., 2007, for a complete description of core project protocols). All study grids were staked and flagged at 30 m intervals, which allowed us to estimate the spatial locations of all squirrel territories to a 10th of an interval, which corresponds to a precision of 3 m. In this study system, territory locations were denoted based on the location of an individual's midden, which approximates the centre of a squirrel's territory. We did not explicitly map territory boundaries for all individuals.

We trapped squirrels using Tomahawk live traps (Tomahawk Live Trap Co., Tomahawk, WI, U.S.A.) baited with peanut butter. If previously tagged, the identities of the squirrels were determined from their unique alphanumeric metal eartags (one in each ear; Monel No. 1; National Band and Tag, Newport, KY, U.S.A.), which they received in their natal nest at around 25 days of age. During the first capture of the season, we marked each squirrel by threading coloured wires through each eartag, which allowed for individual identification of squirrels during behavioural observations. We censused the population twice annually and determined territory ownership through a combination of consistent livecaptures of the same individual at the same midden and behavioural observations of territorial 'rattle' vocalizations (Smith, 1978). Loss of information regarding individual identity was minimal in this system. When squirrels lost an eartag (1.8% of trapping events) they could readily be identified by their remaining tag or by their spatial location on grid (squirrels typically occupy consistent territories throughout their lives; Larsen & Boutin, 1995), and thus could be retagged without loss of information. In extremely rare cases (0.16% of trapping events), squirrels that lost eartags could not be identified and were treated as new individuals.

For each territory owner we defined the social neighbourhood to be all conspecifics whose middens were within a 130 m radius of the owner's midden. One hundred and thirty metres is the farthest distance that red squirrel rattles are known to carry (Smith, 1978) and is similar to the distance at which red squirrels were found to be most responsive to local density changes (150 m; Dantzer et al., 2012), suggesting that 130 m is a reasonable measure of the distance at which red squirrels can receive and respond to acoustic information about their social environment. We measured pairwise familiarity between the territory owner and each neighbour as the number of days that both individuals occupied their current territories within the same acoustic neighbourhood. We then calculated the mean familiarity between the focal individual and all of its neighbours to provide a measure of the average familiarity of each focal individual with its entire acoustic neighbourhood (Siracusa, Boutin, et al., 2017a). We censused the population twice annually, in mid-May and mid-August, because these months correspond to biologically relevant time periods; the May census allowed us to assess the breeding population and the August census allowed us to assess the population at the time of white spruce cone hoarding. We could therefore update each focal squirrel's average familiarity with its neighbours twice per year. This means that we had up to two measures of the social environment for each squirrel in a given vear. So, for example, an individual's average familiarity increased by 90 days if it maintained all of its neighbours from the May census to the August census, and by 270 days if it maintained all of its neighbours from August to May of the following year. Depending on the number of neighbours replaced between censuses, average familiarity could either increase or decrease between these successive time points. If at any point all neighbours were replaced, average familiarity dropped to zero.

Longitudinal Data

Focal observations

Red squirrels are an ideal species for behavioural studies because they are diurnal, easy to locate visually or through acoustic cues, and habituate readily to the presence of humans. As part of the KRSP, we have recorded the behaviour of red squirrels through focal sampling of radiocollared individuals (model PD-2C, 4g, Holohil Systems Ltd, Ontario, Canada) since 1994, although the sampling protocol has varied slightly across this period. In brief, focal animal observations were conducted by using telemetry to locate individuals in the field. Once located, an observer watched the focal animal for a set amount of time (7–10 min, depending on the sampling protocol) and recorded behaviours at 30 s intervals (instantaneous sampling; Altmann, 1974). For this study, we used a subset of long-term behavioural data where focal observations were collected in a consistent manner by instantaneous sampling at 30 s intervals for 10 continuous minutes on a single control grid (SU; N = 487 10-minute sessions over 41 individuals). We excluded any focal observations where the squirrel was out of sight for more than half the observation session (N = 8 10-minute sessions). This provided us with an average of 12 focal observations per individual (range 1-44). These 10 min focal observations were available for female squirrels in 1995 (N = 41), 1996 (N = 10), 1997 (N = 25), 1999 (N = 34), 2001 (N = 70), 2002 (N = 110), 2003 (N = 120) and 2004 (N = 77) and were recorded by 38 different observers between May and August.

Cross-sectional Data

Focal observations

Between 7 May 2016 and 31 August 2016, we used focal animal sampling as described above for seven continuous minutes, rather

than 10, to record red squirrel behaviour (N = 1060 7-minute sessions over 108 individuals). Since rattling is a rare behaviour and is often missed using instantaneous sampling, in 2016 we recorded all occurrences (Altmann, 1974) of rattle vocalizations produced by the focal squirrel, including those which fell outside the 30 s sampling interval (i.e. 'critical incidents'). We used all of these data, including critical incidents, to assess how familiarity affected rattling rates in 2016. Four observers collected behavioural data on both male (N = 76) and female (N = 32) squirrels across two control grids (KL and SU) and one food-supplemented grid (AG). We monitored each individual for 2-10 days consecutively, barring inclement weather (mean = 4 days), and collected an average of 10 focal observations per individual (range 2-29). In instances where multiple focal observations were collected for the same squirrel in a single day, observations were kept 30 min apart at minimum. Because an observer was in regular attendance at these territories, we could be confident that there was no turnover in the social environment during the sampling period for any of these individuals. Territory turnovers in this system are accompanied by substantial rattling and chasing and are therefore easy to detect. The two squirrels for which we observed a disturbance in the local social environment during the sampling period were excluded from this analysis.

For all focal sampling, we recorded and classified red squirrel behaviours in a similar way to previous studies of squirrel behaviour in this system (Anderson & Boutin, 2002; Dantzer et al., 2012; Stuart-Smith & Boutin, 1994). We classified behaviours according to the following categories: vocalizing ('barking' or 'rattling'; Smith, 1978), feeding, foraging, travelling, caching food items, interacting with conspecifics, grooming, resting, vigilant, in nest, or out of sight (unknown behaviour). Vigilance could be distinguished from resting by the alert posture of the squirrel; vigilant squirrels typically had their head up and appeared observant, sometimes standing on their hindlimbs, while resting squirrels often had their head tucked down or lay stretched out.

Audio recording and acoustic analysis

Between 23 June and 25 September 2015 and between 8 May and 1 September 2016, we deployed Zoom H2n audio recorders (Zoom Corporation, Tokyo, Japan) to determine rattling rates of squirrels. We attached recorders with windscreens to 1.5 m stakes and placed a single recorder in the centre of each squirrel's midden. Since Zoom H2n recorders are not weatherproof, we placed an umbrella approximately 30 cm above each audio recorder to protect it from rain and snow. Each morning, we deployed audio recorders between 0500 and 0600 hours (just before squirrels typically became active). We set audio recorders to record in 44.1 kHz/16 bit WAVE format, and recorded in 2-channel surround mode. We allowed audio recorders to run for a full 24 h, but in this study we only use data collected between 0700 and 1300 hours, which is the period during which squirrels are typically most active between early summer and early autumn (Studd, Boutin, McAdam, & Humphries, 2016; Williams et al., 2014). We deployed audio recorders for 137 squirrels (N = 109 males and N = 28 females) and recorded each squirrel for 5 consecutive days on average (range 1–13 days; N = 714 days or 4284 h over 137 individuals). Because we collected audio data over 2 years, we had observations from two to three different social environments for 28 of these individuals, providing some information on within-individual changes in behaviour. Because of the large volume of recordings, we detected rattle vocalizations from recordings automatically using Kaleidoscope software (version 4.3.2; Wildlife Acoustics, Inc., Maynard, MA, U.S.A.). Detection settings included a frequency range of 2000-13 000 Hz, a signal duration of 0.4-15 s, a maximum intersyllable silence of 0.5 s, a fast Fourier transform size of 512 points (corresponding to a frequency resolution of 86 Hz and a temporal resolution of 6.33 ms), and a distance setting of 2 (this value ensures that all detections are retained).

The purpose of using audio recorders was to provide a more accurate estimate of individual rattling rates. One challenge, however, was that the recorders also recorded vocalizations from neighbouring squirrels. Because sound degrades and attenuates predictably with distance, it should be possible to distinguish between the rattles of focal and neighbour squirrels on the basis of rattle acoustic structure. We tested this by conducting hour-long calibrations on 48 focal individuals between 13 September and 14 October 2015. During these calibrations, audio recorders were set up as described above. A single observer standing near the midden kept the territory owner in sight and recorded whether each rattle belonged to the territory owner or a neighbouring individual.

We detected rattle vocalizations from the calibration recordings using Kaleidoscope software (same settings as above). Based on a comparison with the observer's notes, the software detected 100% of the focal squirrel rattles. We then developed a procedure for distinguishing focal squirrel rattles from other types of detections, including neighbour rattles and nonrattles. First, we automatically measured the acoustic structure of every detection using the software package 'Seewave' (version 2.0.5; Sueur, Aubin, & Simonis, 2008) in R (see details of structural measures below). Second, we used the structural measurements in a discriminant function analysis in SPSS (software, version 24; IBM Corporation, Armonk, NY, U.S.A.) to develop a predictive model for assigning detections to groups (i.e. focal rattle, neighbour rattle, nonrattle). We developed the model using detections from half of the 1 h calibration files (selected at random), and then tested it for accuracy by applying it to the detections from the remaining half. The model correctly assigned 80.6% of the focal rattles to the 'focal rattle' group, meaning we missed 19.4% of focal rattles (i.e. false negatives). Some nonrattle detections were also assigned to the 'focal rattle' group, but we removed these by reviewing the spectrograms of all detections categorized as 'focal rattles'. After removing nonrattle detections, 16.0% of the detections remaining in the 'focal rattle' group were false positives, meaning they were actually from the neighbour instead of the focal squirrel. We then applied the predictive model to the main set of audio files, and reviewed all detections labelled as 'focal rattle' in Kaleidoscope to remove the nonrattle detections.

The structural measures included in the discriminant function analysis were (1) duration, (2) root-mean-square amplitude, (3) pulse rate, (4) duty cycle, and five variables that measured the distribution of energy in the frequency domain, including (5) peak frequency, (6) first energy quartile, (7) skewness, (8) centroid and (9) spectral flatness. Duration, root-mean-square amplitude, pulse rate and duty cycle were measured from a waveform. Pulse rate is the number of pulses in the rattle minus one, divided by the period of time between the beginning of the first pulse and the beginning of the last (as in Wilson et al., 2015). Duty cycle is the proportion of the rattle when a pulse is being produced. For pulse rate and duty cycle, individual pulses were identified using the 'timer' function in Seewave (50% amplitude threshold; 200-point smoothing window with 90% overlap). The five energy distribution variables were obtained using the 'specprop' function in Seewave, and were based on a mean frequency spectrum (512-point fast Fourier transform, hanning window, 0% overlap). Peak frequency is the frequency of maximum amplitude. First energy quartile is the frequency below which 25% of the energy is found. Skewness, centroid and kurtosis describe the shape of the power spectrum (detailed definitions can be found in Sueur et al., 2008).

Accelerometers

An accelerometer is an instrument that measures the acceleration of the body along three axes: anterior—posterior (surge), lateral (sway), dorso-ventral (heave) and records temperature, allowing for the characterization of different behavioural patterns. Between 4 May and 1 September 2016, we deployed AXY-3 accelerometers (Technosmart Europe Srl., Rome, Italy) on 94 squirrels (N = 66males, N = 28 females). Accelerometers were deployed in combination with radiotransmitters (model PD-2C, Holohil Systems Ltd, Ontario, Canada). We deployed accelerometers on 94 individuals for an average of 9 days per individual (range 4–17; N = 873 days over 94 individuals) at a sampling rate of 1 Hz. Accelerometers recorded data constantly while deployed, but for this study we only use data between 0600 and 2100 hours to estimate time spent in the nest during active hours of the day (Williams et al., 2014).

Raw accelerometer data were classified into five behavioural categories using threshold values of summary statistics according to the decision tree developed for red squirrel accelerometers and temperature data loggers by Studd et al. (2018). Following methods proposed by Collins et al. (2015), the decision tree was created using 83.8 h of direct behavioural observations on 67 free-ranging squirrels and had an overall accuracy of correctly classifying known behaviours of 94.9% (Studd et al., 2018). Briefly, warm stable temperatures were used to identify when the animal was in the nest with the additional constraint that the individual must not be moving for the majority of each nest bout. Low acceleration values denoted feeding, and high acceleration corresponded to travelling. Travelling was further categorized as running when the peak acceleration value of the surge axis was above a threshold of 1.15 g.

Ethical Note

This study required trapping individuals using Tomahawk live traps in order to attach radiotransmitters and accelerometers. Traps were checked every 60-90 min and squirrels were never left in a trap for longer than 120 min. Radiotransmitters and accelerometers were attached as a single collar around the squirrel's neck using plastic zip ties covered with heat shrink to minimize irritation to the skin. Total package weight for collars with both accelerometers and radiotransmitters (including battery, packaging and bonding material) was 9.6 g on average. For a 200–250 g red squirrel (Steele, 1998) this collar weight was less than the recommended 5-10% of the animal's body weight (Wilson, Cole, Nichols, Rudran, & Foster, 1996). Because all squirrels were continuously monitored through behavioural observations, we could check for irritation caused by the collars. If any irritation was detected (missing fur, red or raw skin around the neck), the squirrel was immediately trapped and the collar removed. Instances of irritation caused by the collar were extremely rare and no squirrels suffered any long-term consequences as a result of the collars. All radiotransmitters and accelerometers were retrieved at the end of the study. Behavioural observations were conducted at least 5 m away from the focal squirrel to minimize any effects of squirrel behaviour and had no detectable negative impact on individuals. This research was approved by the University of Guelph Animal Care Committee (AUP number 1807) and is in compliance with the ASAB/ABS Guidelines for the use of animals in research.

Statistical Analyses

Given that previous work in this study system (Dantzer et al., 2012) allowed us to make specific predictions about how squirrels should adjust rattling rates, vigilance and nest use in response to their social environment, here we used univariate models to test for the effects of familiarity on each of these behaviours explicitly. For all models we included local density, measured as the number of squirrels per hectare within 130 m, as a continuous predictor, to account for the fact that previous work in the study system has found local density to be an important predictor of behavioural time budgets (Dantzer et al., 2012). We also included age as a fixed effect in all rattling rate models since we expected that the vigour of territory defence might decline with physical deterioration, but we did not have specific predictions as to how age might affect nest use or vigilance. Note, however, that since young squirrels are inherently unfamiliar with their neighbours and familiarity increases with age, age and familiarity were correlated (Pearson's correlation coefficient ranged between 0.42 and 0.58 for these analyses), although variance inflation factors were low (<3; Zuur, leno, & Elphick, 2009). Fixed effects and random effects of all models are summarized in Tables 1 and 2, respectively.

Focal data

We analysed the longitudinal (N = 487 10-minute sessions) and cross-sectional (N = 1060 7-minute sessions) focal data separately to account for the structural differences in our data sets. While we had multiple observations of the same individuals across different social environments in the longitudinal data, we only had observations from a single social environment for each of our individuals in the cross-sectional data. In the longitudinal data, there was a single data point where the number of rattles recorded was 25 times greater than the mean. This outlier was likely an error in data entry and was removed (see Appendix, Fig. A1). We analysed the effects of neighbourhood familiarity on (1) the frequency of territorial vocalizations (rattles), (2) the proportion of time spent vigilant and (3) the proportion of time spent in nest. We modelled the frequency of territorial vocalizations using a generalized linear mixed-effect model (GLMM) with a BOBYQA optimizer and a Poisson error distribution (log link) where the response variable was the number of rattles produced during the 10 min focal session. For both the proportion of time spent in nest and the proportion of time spent vigilant, we fitted a beta-binomial model to account for overdispersion in the data (Harrison, 2015). Using the 'cbind' function, we defined the response variable as a two-column matrix composed of the number of observations of the given behaviour (in nest or vigilant) and the number of observations of all other behaviours (not including observations when the squirrel was out of sight). We recognize that the exclusion of observations where the squirrel was out of sight might mean that we are underestimating nest use or vigilance behaviour. However, given that out-of-sight occurrences were relatively rare and only comprise 3% of each focal observation on average, we do not expect this to be a substantial issue. Additionally, if we are underestimating these behaviours, it should be consistent across squirrels and therefore should not bias our results.

In all models we included average familiarity and local density as continuous predictors, and for the rattling rate models we included age as a continuous fixed effect. We included grid, sex and observer identity as categorical fixed effects for the 2016 focal data (it was not necessary to include grid or sex for the longitudinal data as all data were collected on females on a single grid). For both data sets, we included a random intercept term for squirrel identity (squirrel ID) to account for repeated observations of the same squirrels. We wanted to include a random slope term for squirrel ID in the models based on the longitudinal data to assess for individual variation in social plasticity, however, we lacked the statistical power to include this term in our models (Martin, Nussey, Wilson, & Réale, 2011). We did include a random effect of year and observer identity for the longitudinal data set to account for interindividual differences in behavioural scoring.

Audio recorder data

To assess the effects of familiarity on rattling rates derived from the audio recorder data, we fitted a GLMM with a Poisson error

Table 1

Fixed effects from all generalized linear mixed-effects (GLMM) and beta-binomial (BB) models, showing effects of average neighbourhood familiarity, local density and focal squirrel's age on rattling rate, nest use and vigilance behaviour

Method of data collection	Years	Ν	Model	Fixed effect	Parameter \pm SE	Ζ	Р
Longitudinal focals	1995-2004	487					
-			Rattle rate	Familiarity	-0.29 ± 0.12	-2.48	0.01
			(GLMM)	Age	-0.20 ± 0.11	-1.85	0.06
				Density	-0.17 ± 0.13	-1.38	0.17
			Vigilance	Familiarity	0.02 ± 0.13	0.15	0.88
			(BB)	Density	0.05 ± 0.26	0.19	0.85
			Nest use	Familiarity	0.26 ± 0.12	2.31	0.02
			(BB)	Density	-0.37 ± 0.15	-2.53	0.01
Cross-sectional focals	2016	1060					
			Rattle rate	Familiarity	0.07 ± 0.07	1.09	0.27
			(GLMM)	Age	-0.02 ± 0.07	-0.25	0.80
				Density	-0.12 ± 0.06	-2.01	0.04
				Sex-M ^a	-0.13 ± 0.13	-1.02	0.31
				Grid-KL ^b	-0.15 ± 0.13	-1.11	0.27
				Grid-SU ^b	-0.52 ± 0.17	-3.09	0.002
				Obs- JR ^c	0.14 ± 0.25	0.58	0.56
				Obs- MT ^c	-0.34 ± 0.09	-3.87	<0.001
				Obs- YS ^c	-0.47 ± 0.10	-4.92	<0.001
			Vigilance	Familiarity	0.05 ± 0.07	0.69	0.49
			(BB)	Density	-0.01 ± 0.09	-0.07	0.95
				Sex-M ^a	0.26 ± 0.18	1.48	0.14
				Grid-KL ^b	-0.42 ± 0.17	-2.48	0.01
				Grid-SU ^b	-0.36 ± 0.22	-1.63	0.10
				Obs- JR ^c	0.66 ± 0.38	1.73	0.08
				Obs- MT ^c	-0.86 ± 0.16	-5.44	<0.001
				Obs- YS ^c	0.44 ± 0.14	3.24	0.001
			Nest use	Familiarity	-0.11 ± 0.09	-1.21	0.23
			(BB)	Density	0.08 ± 0.10	0.79	0.43
				Sex-M ^a	0.11 ± 0.21	0.52	0.60
				Grid-KL ^b	0.15 ± 0.22	0.70	0.48
				Grid-SU ^D	0.50 ± 0.27	1.85	0.06
				Obs- JR ^c	-0.59 ± 0.57	-1.02	0.31
				Obs- MT ^c	0.41 ± 0.16	2.62	0.009
				Obs- YS ^c	0.37 ± 0.16	2.24	0.02
Audio recordings	2015-2016	714					
			Rattle rate	Familiarity	-0.05 ± 0.04	-1.24	0.21
			(GLMM)	Age	-0.09 ± 0.05	-1.86	0.06
				Density	0.003 ± 0.04	0.08	0.94
				Sex-M ^d	0.01 ± 0.10	0.13	0.90
				Grid-KL ^D	-0.28 ± 0.09	-3.13	0.001
				Grid-SU ⁶	-0.73 ± 0.12	-6.10	<0.001
Accelerometers	2016	873					
			Nest use	Familiarity	-0.0005 ± 0.04	-0.01	0.99
			(BB)	Density	-0.07 ± 0.04	-1.51	0.13
				Sex-M ^d	0.04 ± 0.08	0.53	0.60
				Grid-KL ^D	0.13 ± 0.09	1.36	0.17
				Grid-SU [®]	0.01 ± 0.11	0.11	0.91

Regression coefficients for familiarity, age and density are standardized. Significant effects are indicated in bold.

^a Female taken as the reference.

^b AG (food-supplemented grid) taken as the reference.

^c Observer ES taken as the reference.

distribution (log link). Our response variable was the number of rattles produced between 0700 and 1300 hours (i.e. number of 'focal rattles', unadjusted for false positive or false negative error rates; N = 714 days of recordings). We included average familiarity, local density, age, grid and sex as covariates in the model, as well as a random intercept term for squirrel ID, and an observation-level random effect (OLRE) to account for overdispersion in the model.

Accelerometer data

Using accelerometer data, we assessed the effect of neighbourhood familiarity on the proportion of time spent in nest between 0600 and 2100 hours using a beta-binomial model (N = 873 days). Our response variable was defined as above, using a two-column matrix that included the number of nest observations and the number of observations of all other behaviours. We included average familiarity, local density, grid and sex as fixed effects in the model, and included a random effect for squirrel ID and accelerometer collar.

Exploratory post hoc analysis

Upon finding evidence of behavioural plasticity in the longitudinal data but not the cross-sectional data (see Results below), we conducted an exploratory post hoc analysis in an attempt to understand the inconsistencies in our results. While the longitudinal data provided multiple measures of the same individuals across different social environments, allowing us to estimate withinindividual relationships, our cross-sectional data only allowed us to estimate among-individual relationships. To assess whether our results might be driven by within-individual changes in behaviour, thus limiting our ability to detect behavioural plasticity in the cross-sectional data, we refitted our rattling rate and nest use models from the longitudinal data using a within-subject mean centering approach. Following the methodology of van de Pol and

 Table 2

 Random effects from all generalized linear mixed-effects (GLMM) and beta-binomial (BB) models

Method of data collection	Years	Model	Random effect	Variance	χ^2	df	Р
Longitudinal focals	1995-2004						
-		Rattle rate	Squirrel ID	0.09	2.19	1	0.14
		(GLMM)	Year	< 0.01	< 0.01	1	>0.999
			Observer	0.48	18.55	1	<0.001
		Vigilance	Squirrel ID	0.08	0.56	1	0.45
		(BB)	Year	1.11	5.03	1	0.02
			Observer	0.62	10.53	1	0.001
		Nest use	Squirrel ID	<0.01	<0.01	1	>0.999
		(BB)	Year	0.10	1.20	1	0.27
			Observer	0.10	1.80	1	0.18
Cross-sectional focals	2016						
		Rattle rate (GLMM)	Squirrel ID	0.14	37.53	1	<0.001
		Vigilance (BB)	Squirrel ID	0.14	8.16	1	0.004
		Nest use (BB)	Squirrel ID	0.32	19.20	1	<0.001
Audio recordings	2015-2016						
		Rattle rate	Squirrel ID	0.18	422.38	1	<0.001
		(GLMM)	OLRE	0.06	558.11	1	<0.001
Accelerometers	2016						
		Nest use	Squirrel ID	0.10	157.58	1	<0.001
		(BB)	Accelerometer No.	<0.01	<0.01	1	>0.999

OLRE: observation-level random effect. Significance assessed using a log-likelihood ratio test (LRT) with one degree of freedom to compare models with and without the listed random effect. Significant effects are indicated in bold.

Wright (2009), we split our familiarity term into an amongindividual effect of familiarity (i.e. the mean familiarity score for an individual across all observations) and a within-individual effect of familiarity (i.e. the deviation in each familiarity observation for each individual from their mean score). We applied the same approach to the 2015 and 2016 audio recorder data for which we had some observations from individuals across multiple social environments (Table 3).

Data analysis

We conducted analyses using R version 3.4.1 (R Core Team, 2017) and fitted all GLMMs using the 'Ime4' package (version 1.1–13; Bates, Maechler, Bolker, & Walker, 2015). For all analyses, we fitted generalized additive models to confirm that there were no significant nonlinearities between our predictor and response variables. We checked for overdispersion by comparing the ratio

of the sum of the squared Pearson residuals to the residual degrees of freedom in each model (Zuur et al., 2009) and assessing whether the sum of squared Pearson residuals approximated a chi-square distribution with n - p degrees of freedom (Bolker et al., 2009). As stated above, we accounted for overdispersion in Poisson models by including an observation-level random effect (OLRE; Harrison, 2014). For models with binomial data, we accounted for overdispersion using beta-binomial models, which have been demonstrated to better cope with overdispersion in binomial data (Harrison, 2015). We fitted all beta-binomial models using the package 'glmmADMB' (version 0.8.3.3; Harrison, 2015; Skaug, Fournier, Nielsen, Magnusson, & Bolker, 2018). We standardized all continuous fixed effects to a mean of zero and unit variance. For the following results we present all means \pm SE, unless otherwise stated, and consider differences statistically significant at P < 0.05.

Table 3

Fixed effects from exploratory post hoc models including a within-individual (Familiarity_w) and among-individual (Familiarity_A) effect of familiarity

Method of data collection	Years	Ν	Model	Fixed effect	Parameter \pm SE	Z	Р
Longitudinal focals	1995-2004	487					
-			Rattle rate	Familiarity _W	-0.21 ± 0.08	-2.51	0.01
			(GLMM)	Familiarity _A	-0.18 ± 0.12	-1.50	0.13
				Age	- 0.22 ± 0.11	-2.02	0.04
				Density	-0.18 ± 0.13	-1.42	0.15
			Nest use	Familiarity _W	0.17 ± 0.10	1.68	0.09
			(BB)	Familiarity _A	0.19 ± 0.13	1.53	0.13
				Density	-0.37 ± 0.15	-2.50	0.01
Audio recordings	2015-2016	714					
			Rattle rate	Familiarity _W	-0.03 ± 0.01	-2.55	0.01
			(GLMM)	Familiarity _A	0.02 ± 0.05	0.34	0.74
				Age	-0.10 ± 0.05	- 2.09	0.04
				Density	-0.02 ± 0.04	-0.56	0.58
				Sex-M ^a	0.01 ± 0.10	0.12	0.90
				Grid-KL ^b	-0.28 ± 0.09	-3.16	0.002
				Grid-SU ^b	-0.81 ± 0.12	-6.58	<0.001

GLMM: generalized linear mixed-effects models; BB: beta-binomial models. Regression coefficients for familiarity, age and density are standardized. Significant effects are indicated in bold.

^a Female taken as the reference.

^b AG (food-supplemented grid) taken as the reference.

RESULTS

Heterogeneity in mean neighbourhood density and familiarity were very similar between the longitudinal and cross-sectional data sets. Among the years in which we analysed long-term focal data (1995–2004), variation in average neighbourhood familiarity ranged from 0 (corresponding to when a squirrel first established its territory) to 813 days (mean: 229 ± 9 days) and variation in local density ranged from 0.57 to 5.84 squirrels/ha (mean: 1.93 ± 0.05 squirrels/ha). In our 2015 and 2016 data, there was a nearly equivalent amount of variation in average neighbourhood familiarity and local density. Neighbourhood familiarity ranged from 0 to 855 days (mean: 296 ± 5 days) and local density ranged from 1.13 to 6.03 (mean: 3.34 ± 0.03 squirrels/ha). Below we discuss the effects of familiarity and age on behavioural patterns. Results for other fixed effects in the models can be found in Table 1; random effects for all models can be found in Table 2.

Longitudinal Data

Territorial defence

During the long-term focal observations, red squirrels produced an average of 0.37 ± 0.04 rattles per 10 min observation session (range 0–4), which is equivalent to one rattle every 27.06 min. Red squirrels in the longitudinal data set adjusted their behaviour in response to increasing average neighbourhood familiarity by producing significantly fewer rattles ($\beta = -0.29 \pm 0.12$, z = -2.48, P = 0.01; Fig. 1). This corresponds to a predicted three-fold decrease in rattling rates: in neighbourhoods with the lowest familiarity, squirrels were predicted to rattle once every 24.76 min and in neighbourhoods with the highest familiarity, only once every 79.75 min. The effect of age on rattling rates was marginally nonsignificant ($\beta = -0.20 \pm 0.11$, z = -1.85, P = 0.06; Table 1). On average, squirrels spent $6.0 \pm 0.7\%$ of their time vigilant, but did not show changes in vigilance behaviour in response to changing familiarity with neighbours ($\beta = 0.02 \pm 0.13$, z = 0.15, P = 0.88; Table 1).

Nest use

Based on the longitudinal data, red squirrels spent, on average, $31.0 \pm 2.0\%$ of their time in nest. Red squirrels responded to changing social conditions by increasing nest use in response to increasing familiarity ($\beta = 0.26 \pm 0.12$, z = 2.31, P = 0.02; Fig. 1). This is equivalent to a predicted 23% increase in nest use: squirrels in neighbourhoods with the lowest familiarity were predicted to spend only 22% of their time in nest compared to 45% in neighbourhoods with the highest familiarity.

If we bin the data for both these analyses, and use binomial models to look at the probability of producing at least one rattle, or spending more than 50% of time in the nest, the effects weaken slightly, as we expect would occur when collapsing variation in the data, but the patterns remain the same (see Appendix, Fig. A2).

Cross-sectional Data

Territorial defence

During focal observations in 2016, red squirrels produced 0.71 ± 0.03 rattles per 7 min observation session (range 0-6), which equates to approximately one rattle every 9.80 min. Data from audio recorders in 2015 and 2016 provided very similar estimates of rattling rates. We captured, on average, 33.96 ± 0.72 rattles per 6 h of recording (range 3–123), which, after correcting for the error rates in our discriminant function analysis, is equivalent to one rattle every 9.81 min. Rattling rates were much higher than in the longitudinal data due to differences in behavioural sampling protocol. In 2016, all occurrences of rattling were recorded as 'critical incidents', while in the longitudinal data, rattles were only recorded if they fell on the 30 s sampling interval. When critical incidents of rattling were removed from the 2016 data, rattling rates dropped to one rattle every 40.55 min. Based on both crosssectional focal observations and audio recorder data, neither average familiarity of the social neighbourhood (all |z| < 1.25, all P > 0.21) nor age (all |z| < 1.87, all P > 0.06) were significant predictors of rattling rate (Table 1). Focal observations indicated that



Figure 1. Red squirrels adjust (a) rattling rate and (b) proportion of time spent in nest in response to the average familiarity of their social neighbourhood (*N* = 487). Results are based on 10 min behavioural observations of squirrels during 1995–2004. Values on the *X* axis are standardized measures of average familiarity. Points indicate raw data with a small amount of jitter introduced to show overlapping points.

red squirrels spent 7.0 \pm 0.5% of their time vigilant, on average, but did not adjust vigilance behaviour in response to changing familiarity with neighbours ($\beta = 0.05 \pm 0.07$, z = 0.69, P = 0.49; Table 1).

Nest use

Based on focal observations in 2016, red squirrels spent an average of $36.0 \pm 1.0\%$ of their time in nest. Accelerometer data from 2016 provided similar estimates of the average proportion of time spent in nest during daylight hours ($36.0 \pm 0.4\%$). Both focal observations and accelerometer data indicated that squirrels did not adjust their nest use in response to familiarity with neighbours (all |z| < 1.22, all P > 0.22; Table 1).

Exploratory post hoc analysis

In our post hoc analyses we found evidence to suggest that effects of familiarity on rattling rates were primarily due to withinindividual changes in behaviour rather than among-individual differences. In the longitudinal data, increasing familiarity led to a significant decrease in rattling rates within ($\beta = -0.21 \pm 0.08$, z = -2.51, P = 0.01) but not among individuals ($\beta = -0.18 \pm 0.12$, z = -1.50, P = 0.13; Table 3). There were positive within- and among-individual effects of familiarity on nest use, but neither of these effects were significant (all |z| < 1.69, all P > 0.08; Table 3). Audio recorder data from 2015 and 2016 also a revealed a significant negative within-individual effect of familiarity on rattling rates ($\beta = -0.03 \pm 0.01$, z = -2.55, P = 0.01), but no among-individual effect ($\beta = 0.02 \pm 0.05$, z = 0.34, P = 0.74; Table 3). Results from the audio data should be interpreted with caution as the inclusion of year in the model affected these results (see Appendix, Table A1).

DISCUSSION

For territorial species, the ability to be responsive to changes in the social environment may convey a fitness advantage by allowing individuals to reduce time and energy investment in costly behaviours (Höbel, 2015; Krobath, Römer, & Hartbauer, 2017; Ydenberg, Giraldeau, & Falls, 1988). In this study, we used multiple types of behavioural data, as well as a longitudinal and crosssectional data set, to test a single overarching hypothesis: that red squirrels show behavioural plasticity in response to the familiarity of their social neighbourhood. Our results provide evidence that solitary North American red squirrels can respond to changes in the composition of their social environment and that they do so under natural conditions and in a manner that is consistent with our expectations for adaptive behavioural change in this species. Although our evidence for social plasticity comes exclusively from female squirrels, both male and female red squirrels defend exclusive territories based around a central cache of food resources and produce territorial vocalizations that are the same in both form and function (Smith, 1968). Given this, and the fact that we found no evidence of an interaction between familiarity and sex in the cross-sectional analyses (Appendix, Table A2), we have no reason to expect that social plasticity differs between male and female red squirrels.

Previous work in this study system has demonstrated that red squirrels face reduced intrusion risk in social neighbourhoods with high average familiarity (Siracusa, Boutin, et al., 2017a). As such, we predicted that red squirrels would show appropriate social plasticity by reducing territorial defence behaviours and increasing time and energy spent on self-maintenance behaviours when familiar with neighbouring conspecifics. Results from behavioural observations across 8 years provided support for these predictions, indicating that red squirrels demonstrated social plasticity by reducing rattling rates and increasing the proportion of time spent in nest in social neighbourhoods with high average familiarity

(Fig. 1). Such changes in behaviour not only minimize the time spent on territory defence but might also reduce associated costs of territoriality. Territorial vocalizations may attract the attention of predators (Abbey-Lee, Kaiser, Mouchet, & Dingemanse, 2016), and rattles are loud, broadband signals that should be easy to localize (Marler, 1955). By reducing rattling rates under less risky social conditions, squirrels may benefit from reduced predation risk. Additionally, spending more time in the nest when familiarity with neighbours is high also presumably reduces the risk of being detected by a predator.

We did not, however, find effects of neighbourhood familiarity on vigilance behaviour. This could be due to vigilance for conspecifics being easily confounded with vigilance for predators. In contrast, Dantzer et al. (2012) found significant effects of local density on vigilance using behavioural data collected over a similar time frame, indicating that conspecific rather than heterospecific effects on vigilance are detectable in this study system. While we included local density as a covariate in all of our models to account for the potential effects of density on behaviour (Dantzer et al., 2012), our goal was not to directly estimate effects of density, and our results therefore are not a clean representation of density effects. In several cases density was correlated with other variables in the model, such as grid, leading to substantial changes in the parameter estimates for density. As a result, the effects of density on behaviour that we report here cannot be compared directly to previous studies of these effects in this population (e.g. Dantzer et al., 2012; Shonfield, Taylor, Boutin, Humphries, & McAdam, 2012) and we do not discuss the effects of density further.

Results from the cross-sectional data in 2015 and 2016 did not corroborate our longitudinal results showing behavioural responses to familiarity. Findings from the focal observations, audio recorders and accelerometers indicated that when using amongindividual relationships to estimate the effects of the social environment on behaviour, there was no effect of familiarity on territorial behaviours (rattling rates, vigilance) or self-maintenance (nest use; Table 1). While these results were surprising, such inconsistencies between longitudinal and cross-sectional results are well documented in both the sociological (Chassin, Presson, Sherman, & Edwards, 1992; Easterlin, 1974) and statistical literature (Simpson, 1951). In the present study, the inconsistencies we observed were most likely because analysis of the cross-sectional data was largely among individuals. Since behavioural plasticity is functionally a within-individual phenomenon, using amongindividual differences in behaviour to estimate plasticity relies on the assumption that the among-individual relationship is an accurate representation of within-individual changes in behaviour. Using a within-subject centering approach (van de Pol & Wright, 2009), we found that the within- and among-individual effects were not equivalent. In the longitudinal data, we found that individuals adjusted rattling rates in response to changes within their own social environment (i.e. a significant within-individual effect), but we did not observe significant differences in rattling rates when comparing among individuals (Table 3). Similarly, for the audio recorder data (the only cross-sectional data for which we had some observations of individuals across multiple social environments), we found evidence of a significant within-individual, but not among-individual, effect (Table 3). Thus, while we clearly see evidence of plasticity when considering changes in individual behaviour across different social environments, in this study system it appears that we cannot estimate these effects by comparing behaviour among individuals.

One potential explanation for this discrepancy is that the among-individual effect is masked by individual variation in plasticity, whereby substantially different individual 'slopes' result in a 'mean slope' of zero (i.e. the absence of a significant population-



Figure 2. Three different scenarios for how variation in mean rattling rate (random intercepts) in combination with variation in data sampling structure might change our ability to detect among-individual effects when individuals have the same slope. We schematically depict the within-individual slopes (solid grey lines) of seven subjects (j = 1 to j = 7). The solid grey lines indicate the range over which each individual was sampled. Dotted lines provide an extension of these slopes to the edge of the figure. The among-subject slope (solid black line) is based on the association between \bar{x}_i and \bar{y}_i as denoted by the filled black circles.

level response to the environment; Nussey et al., 2007). We were unable to test this hypothesis as we lacked the statistical power to include a random slope term in our models (Martin et al., 2011). Furthermore, even if all individuals demonstrate negative reaction norms (i.e. reduced rattling rate in response to increasing familiarity), there are still several reasons we might fail to detect differences among individuals. First, it seems unlikely that squirrels can assess their absolute familiarity, meaning that behavioural adjustments are dependent on the relative social environments individuals experience rather than absolute changes in familiarity. Nor were we able to precisely measure absolute familiarity since measures of the social environment were based on semiannual census data, which may have added some noise to the data. Additionally, variation in individual mean ratting rates (i.e. random intercepts) due to differences in sex, age, personality, stress, among other possibilities, might mask an among-individual effect. These factors, combined with variation in the range of social environments sampled for a given individual, mean that, even when all individuals show negative reaction norms, it is possible to measure a lack of (Fig. 2b), or even a positive among-individual effect (Fig. 2c). Additional individual data, spanning a range of social environments, is necessary to better understand the patterns leading to within- versus among-individual effects in this system.

While measuring the same individuals across multiple environments is the cleanest way to assess behavioural plasticity, obtaining such measures is often extremely time and labour intensive and may not even be possible on short timescales. In such instances, studies may use measures from unique individuals in different environments to approximate this plasticity (see Legagneux & Ducatez, 2013; Slabbekoorn & Peet, 2003, for examples). Importantly, such studies unavoidably confound environmental differences with individual identity. Therefore, in instances where behavioural differences among environments are found, such differences cannot be exclusively attributed to plasticity but may instead be due to genetic or other consistent differences among individuals. In addition, even in instances where behavioural differences among individuals are not present, our results suggest that studies should be cautious about interpreting this as a lack of plasticity since the estimated among-individual effect may not be an accurate reflection of withinindividual changes in behaviour.

Although we have provided an explanation for the differences in our longitudinal and cross-sectional findings, there are a couple reasons why it is important to interpret our results with caution. First, there is potential for changes in rattling rates to be driven by effects of age rather than familiarity if the strength of territory defence declines with physical deterioration. This type of linear senescent decline is evident in other traits in red squirrels (Descamps, Boutin, Berteaux, & Gaillard, 2008; McAdam et al., 2007). We have done our best to account for this possibility in our analyses, but given that these variables are strongly correlated, an experimental approach would prove useful in disentangling these effects, as they are difficult to tease apart statistically. Second, it is worth addressing our use of multiple univariate analyses to test a single overarching hypothesis. Previous research in this study system has detected effects of the social environment on vigilance and nest use using a multivariate analysis (Dantzer et al., 2012), allowing us to make specific predictions about how squirrels should adjust patterns of nest use and vigilance in response to neighbourhood familiarity. Given this, we felt that analysing the effects of familiarity on each of these behaviours individually provided a more elegant test of our hypothesis. However, our use of univariate analyses increases our chances of committing a type I error by attributing variance as unique to a single response variable when it may in fact be shared (Huberty & Morris, 1989).

Despite these limitations, we believe that the results from our study, in particular the data for which we can estimate withinindividual changes in behaviour, provide evidence that red squirrels are socially plastic. Furthermore, although we have not directly tested the fitness consequences of social plasticity, red squirrels reduced rattling rates, thereby spending less time on territory defence and potentially minimizing risk of detection by predators, under social conditions where intrusion risk was low (Siracusa, Boutin, et al., 2017a). This suggests that 'asocial' species can not only be socially responsive but also socially competent in their behaviour (Taborsky & Oliveira, 2013, 2012). While evidence for reduced aggression towards familiar conspecifics is taxonomically widespread (reviewed in Temeles, 1994), these studies have typically been focused on documenting behavioural changes on short timescales through exposure to an experimental stimulus. Our study demonstrates that natural variation in neighbourhood familiarity has direct consequences for behavioural time budgets by allowing individuals with familiar neighbours to reduce territory defence and increase time spent in nest. Only a handful of previous studies have demonstrated similar patterns in wild populations under natural social conditions. Willow ptarmigan, Lagopus lagopus, males were found to spend significantly more time engaged in territorial border disputes when they had more new neighbours (Eason & Hannon, 1994). In Seychelles warblers, Acrocephalus

sechellensis, living near familiar individuals provided important benefits by reducing immediate energetic costs through fewer physical fights (Bebbington et al., 2017).

Additionally, recent research has increasingly noted the importance of group composition in shaping individual behaviour (Farine, Montiglio, & Spiegel, 2015). For example, nutmeg mannikins, Lonchura punctulata, have been shown to forgo consistent individual differences in scrounger-forager tactics when flock composition changes, and to adjust their social strategy according to frequencydependent payoffs (Morand-Ferron, Wu, & Giraldeau, 2011). Water striders (Aquarius remigis) also show plasticity in aggression and activity in response to the presence of hyperaggressive individuals in the group (Sih et al., 2014) or changes in male-male competition (Montiglio et al., 2017). Although territorial species do not act in clearly defined, discrete units, we have demonstrated that red squirrels show similar social plasticity in response to the composition of neighbouring territory holders at the scale of the acoustic social environment (i.e. 130 m radius). Our results emphasize that the composition of neighbouring conspecifics, in addition to the quantity of individuals in the social environment (Dantzer et al., 2012), can shape the behaviour of territorial species.

Conclusion

It has been recognized for decades that familiarity with neighbours may help to reduce the costs of territorial conflict. However, the importance of these social relationships for mitigating time spent on defence under natural conditions has rarely been explored. Here we show that free-living red squirrels minimize costs of defence by reducing rattling rates three-fold and increasing nest use by approximately 25% when familiarity with neighbours is high and intrusion risk is therefore low. Taken together, these results provide evidence that solitary, territorial species have the capacity to assess and respond to nuanced changes in their social environment, despite not typically being considered to engage in important social interactions. Importantly, our results also suggest that behavioural plasticity in this species cannot be estimated by comparing differences in behaviour among individuals, emphasizing the need to have observations from the same individuals across multiple social environments in order to detect these behavioural patterns.

Data Availability

Data are available from the Figshare Repository (https://doi.org/ 10.6084/m9.figshare.7866380).

Acknowledgments

We thank the three anonymous referees who provided constructive feedback on previous versions of this manuscript. We are grateful to all of the field technicians who have contributed to the long-term KRSP database over the years. We are particularly indebted to M. Thorpe, Y. Sun and J. Robertson for their endless hours of squirrel watching and for waking up at unreasonable hours of the morning to deploy audio recorders. We acknowledge that this study was conducted on Champagne and Aishihik First Nations land and thank Agnes MacDonald and her family for longterm access to her trapline. This research was supported by Discovery Grants and Northern Research Supplements from the Natural Sciences and Engineering Research Council of Canada (A. G. McAdam: RGPIN-2015-04707); S. Boutin: RGPIN-2014-05874; D. R. Wilson: RGPIN-2015-03769; J. E. Lane: RGPIN-2014-04093, RGPNS-2014-459038, as well as funding from the University of Michigan (B. Dantzer), the W. Garfield Weston Foundation (E. K. Studd) and Grants-in-Aid of research from the American Society of Mammalogists and the Arctic Institute of North America (E. R. Siracusa). This is publication number 95 of the Kluane Red Squirrel Project.

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Appendix



Figure A1. Distribution of the number of rattles produced per 10 min focal observation. Grey bars show raw data (N = 488 observations); black outline indicates a theoretical Poisson distribution simulated using N = 500 data points. The single outlier (where 9 rattles were recorded in a single focal observation) was removed before data analysis, as it appeared to be a data entry error.



Figure A2. Binomial models looking at the effect of familiarity with neighbours on the probability of (a) producing at least one rattle ($\beta = -0.23 \pm 0.15$, z = -1.52, P = 0.13) and (b) spending more than 50% of time in nest ($\beta = 0.25 \pm 0.12$, z = 2.05, P = 0.04). These models offer an additional way of analysing the longitudinal focal data presented in this manuscript. However, we note that by binning the data into categories we are inherently collapsing variation in the data and thereby reducing our power to detect an effect of familiarity on behavioural patterns. It is therefore unsurprising that in these models the effects of familiarity slightly weaken but the patterns remain the same.

Table A1

Results from the audio recorder data showing effects of within-individual (Familiarity_w) and among-individual (Familiarity_A) familiarity on rattling rate when including year as a fixed effect

Model	Fixed effect	Parameter \pm SE	Z	Р
Rattle rate (GLMM)	Familiarity _W Familiarity _A Age Density Sex-M ^a Grid-KL ^b Grid-SU ^b Year-2016	$\begin{array}{c} 0.02 \pm 0.02 \\ 0.05 \pm 0.04 \\ -0.09 \pm 0.04 \\ -0.09 \pm 0.04 \\ -0.15 \pm 0.10 \\ -0.23 \pm 0.08 \\ -0.75 \pm 0.12 \\ -0.40 \pm 0.08 \end{array}$	1.30 1.17 -1.93 - 2.25 -1.53 - 2.76 - 6.40 - 5.13	0.19 0.24 0.05 0.02 0.13 0.006 <0.001 <0.001

The inclusion of year helps to account for structure in our data, however, we had no a priori hypothesis for why year itself might affect behavioural patterns. Our expectation was that changes in rattling rates would be driven by changes in density or familiarity between years. Given this, and the fact that year was correlated with familiarity (r = 0.34) and density (r = -0.43), we excluded year from our primary analysis. We have included year here to be transparent about its effects in the model. Significant effects are indicated in bold.

^a Female taken as the reference.

^b AG (food-supplemented grid) taken as the reference.

Table A2 Interaction between familiarity and sex from all cross-sectional generalized linear mixed-effects (GLMM) and beta-binomial (BB) models

Method of data collection	Years	Ν	Model	Fixed effect	Parameter \pm SE	Z	Р
Cross-sectional focals	2016	1060					
			Rattle rate (GLMM)	Familiarity * sex	-0.04 ± 0.12	-0.31	0.75
			Vigilance (BB)	Familiarity * sex	-0.02 ± 0.17	-0.14	0.89
			Nest use (BB)	Familiarity * sex	0.12 ± 0.20	0.61	0.54
Audio recordings	2015-2016	714					
			Rattle rate (GLMM)	Familiarity * sex	-0.11 ± 0.09	-1.15	0.25
Accelerometers	2016	873					
			Nest use (BB)	Familiarity * sex	-0.03 ± 0.08	-0.31	0.76

Models include the same covariates as shown in Table 1, but for the sake of simplicity only the interaction term is shown here.