

Unit 9

Probability

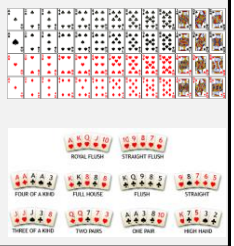
Computer Science 1002

Introduction to Logic for Computer Scientists

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Puzzle: playing poker


- There are 52 cards in a standard deck; 4 suits of 13 ranks each.
- In poker, some 5-card combinations ("hands") are special:
 - For example, a "three of a kind" consists of three cards with the same rank, together with two arbitrary cards.
- What are the chances to get
 - a three of a kind hand?
 - A two pairs hand (5 cards with 2 same-rank pairs)?
 - Other hands?...



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Finite probability

- Experiments:** producing an **outcome** out of possible choices
 - Tossing a coin: outcome can be "heads"
 - Getting a lottery ticket: outcome can be "win"
- Sample space S:** set of all possible outcomes.
 - {heads, tails} for a coin toss
 - $\{1,2,3,4,5,6\} \times \{1,2,3,4,5,6\}$ for rolling two dice
- Event $A \subseteq S$:** subset of outcomes
 - Both dice came up even.




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Finite probability

- Probability** of an event if all outcomes are **equally likely**:
 - $\Pr(A) = |A|/|S|$ (fraction of the outcomes that are in the event A).
 - Probability of both dice coming up even:
 - $A = \{(2,2), (2,4), (4,2), (2,6), (6,2), (4,4), (4,6), (6,4), (6,6)\}$.
 - $|A| = 9, |S| = 36$
 - $\Pr(A) = 9/36 = 1/4$
- Probability is always a number between 0 (event cannot happen, $A = \emptyset$) and 1 (event always happens, $A = S$).
- If $\Pr(A) = p$, then $\Pr(\bar{A}) = 1 - p$ (that is, probability that A does not happen is 1 – probability of A.)


Let's use combinatorics we just studied to calculate probabilities!



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Puzzle: playing poker


- There are 52 cards in a standard deck; 4 suits of 13 ranks each.
- In poker, some 5-card combinations ("hands") are special:
 - For example, a "three of a kind" consists of three cards with the same rank, together with two cards of other different ranks.
- How many ways are there to choose (ignoring the order)
 - A royal flush?
 - a three of a kind hand?
 - a two pairs hand?
 - other hands?...



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Puzzle: playing poker

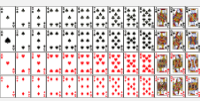

- There are 52 cards in a standard deck; 4 suits of 13 ranks each.
- How many ways are there to choose a royal flush
 - all same suit, ranks A, K, Q, J, 10.
 - Pick a suit: $C(4,1) = 4$
- How many ways to choose three of a kind?
 - pick the rank: $13 = C(13,1)$
 - Pick 3 out of 4 kinds of this rank: $4 = C(4,3)$
 - Pick two other ranks: $C(12,2) = 66$
 - Pick a suit of each of the other ranks: $C(4,1) \cdot C(4,1) = 16$
 - Total: $13 \cdot 4 \cdot 66 \cdot 16 = 54912$



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Puzzle: playing poker

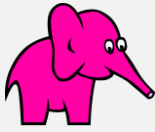

- What is the probability of getting a three of a kind hand?
 - Experiment: drawing a hand (5 cards).
 - Sample space S: set of all possible hands.
 - Size of S is $C(52, 5) = \binom{52}{5} = 2,598,962$
 - Event A: getting a three of a kind hand
 - Size of the event A: 54,912
 - Probability of A (all hands are equally likely):
 - $\Pr(A) = \frac{|A|}{|S|} = 0.0211..$

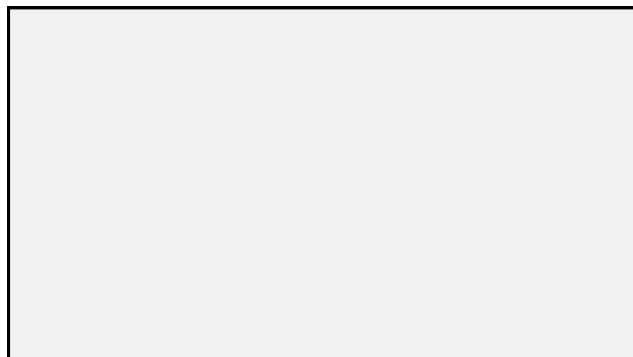
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Probabilities and pink elephants

- What is the probability that walking down George street you'd see a pink elephant?
 - Your friend says: "It is $\frac{1}{2}$! You will either see the pink elephant, or not!"
 - Do you agree?

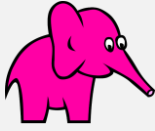

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Probabilities and distributions



- What if outcomes are not equally likely?
 - Biased coins, pink elephants, etc.
- A function $\Pr: S \rightarrow \mathbb{R}$ is a **probability distribution** on (a finite set) S if \Pr satisfies the following:
 - For any outcome $s \in S$, $0 \leq \Pr(s) \leq 1$
 - $\sum_{s \in S} \Pr(s) = 1$
- Uniform distribution:** for all $s \in S$, $\Pr(s) = 1/|S|$
 - all outcomes are equally likely
 - Fair coin: $\Pr(heads) = \Pr(tails) = \frac{1}{2}$


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Biased coin

- Lots of distributions other than uniform!
- Biased coin: say heads twice as likely as tails.
 - $\Pr(heads) + \Pr(tails) = 1$
 - $\Pr(heads) = 2 * \Pr(tails)$
 - So $\Pr(heads) = \frac{2}{3}$, $\Pr(tails) = \frac{1}{3}$

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Probabilities of events


- Probability of an event A is a sum of probabilities of the outcomes in A

$$\Pr(A) = \sum_{a \in A} \Pr(a)$$
 - Probability of A not occurring: $\Pr(\bar{A}) = 1 - \Pr(A)$
- Probability of the union of two events (either A or B happens) is

$$\Pr(A \cup B) = \Pr(A) + \Pr(B) - \Pr(A \cap B)$$
 - By principle of inclusion-exclusion
 - If A and B are disjoint, $\Pr(A \cap B) = 0$, so $\Pr(A \cup B) = \Pr(A) + \Pr(B)$
- In general, if events $A_1 \dots A_n$ are **pairwise disjoint**
 - that is, $\forall i, j$ if $i \neq j$ then $A_i \cap A_j = \emptyset$
 - Then $\Pr(\bigcup_{i=1}^n A_i) = \Pr(A_1 \cup A_2 \cup \dots \cup A_n) = \sum_{i=1}^n \Pr(A_i)$




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Probabilities of events

- Suppose a die is biased so that 3 appears twice as often as any other number (others equally likely).
 - Probability of 3: $2/7$. Probabilities of others: $1/7$
- What is the probability that an odd number appears?
 - Event: $A = \{1, 3, 5\}$
 - $\Pr(A) = 1/7 + 2/7 + 1/7 = 4/7$.
- What is a probability that an odd number or a number divisible by 3 appears?
 - $A = \{1, 3, 5\}$, $B = \{3, 6\}$, $A \cap B = \{3\}$
 - $\Pr(A) = 4/7$, $\Pr(B) = 3/7$, $\Pr(A \cap B) = 2/7$
 - $\Pr(A \cup B) = \Pr(\{1, 3, 5, 6\}) = \frac{4}{7} + \frac{3}{7} - \frac{2}{7} = \frac{1}{7} + \frac{2}{7} + \frac{1}{7} + \frac{1}{7} = \frac{5}{7}$




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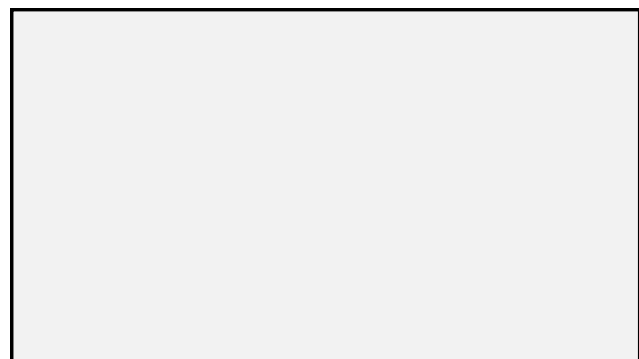


Birthday paradox


- How many people have to be in the room so that probability that two of them have the same birthday is at least $\frac{1}{2}$?



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Birthday paradox

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Birthday paradox

- How many people have to be in the room so that probability that two of them have the same birthday is at least $\frac{1}{2}$?
 - Considering all birthdays independent: no twins!
 - If there are twins in the room, then the probability is 1.
 - So no twins is the harder case.
 - And considering all days equally likely
 - Otherwise again probability would be higher.
 - Even counting leap years: 366 days.
 - So each person has probability $1/366$ to have a birthday on any given day.

* The real people have to be in the room so that probability that two of them have the same birthday is at least 1/2
- Counting all birthdays independent: no twins!
- Otherwise again probability would be higher.
- Accounting for leap years
- Leap years are not equally likely
- Leap years are not equally likely
- Leap years are not equally likely

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Birthday paradox

- Let's look at the first two people, first three, etc.
 - And calculate the probabilities of them all having different birthdays.
- Product rule: number of combinations of distinct birthdays of the first i people is $P(366, i) = 366 \cdot 365 \cdot \dots \cdot (366 - i + 1)$
 - Probability that the first i people all have different birthday is $\frac{P(366, i)}{366^i} = \frac{365 \cdot 364 \cdot \dots \cdot (366 - i + 1)}{366 \cdot 366 \cdot \dots \cdot 366}$
 - So with probability $1 - \frac{P(366, i)}{366^i}$ at least two out of first i people have birthday on the same day.
 - That works out to about $i = 23$ people to reach $\frac{1}{2}$.

Only 23! That's why it is called a paradox.



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Puzzle: Monty Hall problem

- Let's make a deal!
 - A player picks a door.
 - Behind one door is a car.
 - Behind two others are goats.
- A player chooses a door.
 - A host opens another door
 - Shows a goat behind it.
 - And asks the player if she wants to change her choice.
- Should she switch?



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Puzzle: Monty Hall problem

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 - A player has to pick a door.
 - Behind one door is a car.
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Puzzle: Monty Hall problem

- A player has to pick a door. Behind one door is a car. Behind other doors are goats.
- After the player chooses a door, the host opens another door, shows a goat behind it, and asks the player if she wants to change her choice.
- Should she switch?
 - Originally, probability of picking the car is $\frac{1}{3}$
 - If she first picked a door with a car: ($\frac{1}{3}$ probability)
 - Then she would switch to a goat.
 - If she first picked a door with a goat ($\frac{2}{3}$ probability)
 - Then she would switch to a car.
 - Yes, she should switch to increase her probability of getting a car!




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Conditional probabilities



- Conditional probability** of an event A given event B, denoted $\Pr(A|B)$, is the probability of A if we know that B occurred.
 - Probability of rain given that it is cloudy is higher than just probability of rain.
 - Probability of rain is lower given that it is sunny.
 - In the Monty Hall puzzle
 - Probability of car behind door 2 is $\frac{1}{3}$
 - Probability of a car behind door 2 conditional the host showing a goat behind door 3 is $\frac{2}{3}$
- It is probability of both A and B, given that B happened for sure:

$$\Pr(A|B) = \frac{\Pr(A \cap B)}{\Pr(B)}$$
 - Assume that $\Pr(B) > 0$: after all, B did happen.

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Conditional probabilities


- Suppose that there are 9,000 black bears and 1000 polar bears in our province, NL. Moreover, all polar bears and 2000 of black bears are in Labrador.
 - Let A be an event that a bear is a polar bear.
 - Let B be an event that a bear is in Labrador (rather than Newfoundland).
- Probability that a random bear in NL is a polar bear:

$$\Pr(A) = \frac{1000}{1000+9000} = \frac{1}{10}$$
- Probability that a random bear in NL is in Labrador:


$$\Pr(B) = \frac{1000+2000}{1000+9000} = \frac{3}{10}$$
- Probability of being both a polar bear and in Labrador is $\Pr(A \cap B) = \frac{1}{10}$
 - Same as probability of being a polar bear, since all polar bears are in Labrador
- Probability that a random bear is a polar bear conditional on it being in Labrador:

$$\Pr(A|B) = \frac{\Pr(A \cap B)}{\Pr(B)} = \frac{1/10}{3/10} = \frac{1}{3} = \frac{1000}{1000+2000}$$
 - With this data $\Pr(A|\bar{B}) = 0$. A random bear in Newfoundland is a black bear.

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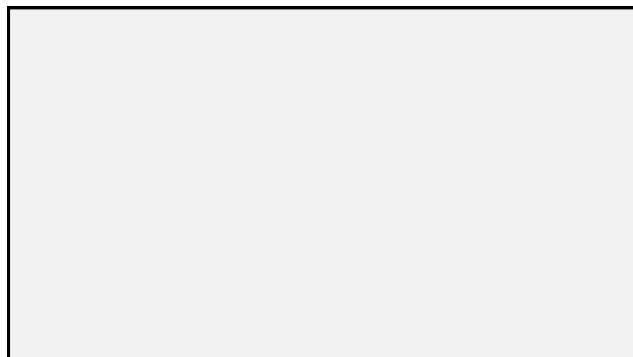


Independent events




- If knowing B gives us no information about A and vice versa, then A and B are **independent** events:
 - Then $\Pr(A) = \Pr(A|B) = \frac{\Pr(A \cap B)}{\Pr(B)}$.
 - So A and B are independent iff $\Pr(A \cap B) = \Pr(A) \cdot \Pr(B)$.
 - In general, events $A_1 \dots A_n$ can be **pairwise independent** (that is, any two A_i, A_j are independent, or (stronger condition) **mutually independent**: $\forall T \subseteq \{A_1, \dots, A_n\} \quad \Pr(\bigcap_{A_i \in T} A_i) = \prod_{A_i \in T} \Pr(A_i)$)
 - That is, for every subset of these events, probability of them occurring together is the product of individual probabilities.
 - Different coin tosses/dice rolls are usually considered independent.

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Contrapositive vs. Converse



- "If a person is carrying a weapon, then the airport metal detector will ring".
 - Same as "If the airport metal detector does not ring, then the person is not carrying a weapon".
 - Not the same as: "If the airport metal detector rings, then the person is carrying a weapon."
- "If the person is sick, then the test is positive".
- "If he is a murderer, his fingerprints are on the knife".

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Sensitivity and specificity



- Let A: "metal detector rings", B: "person carries a weapon",

Sensitivity:
percentage of correct **positives**


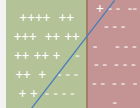
- $\Pr(A|B) = \Pr(A \cap B) / \Pr(B)$
- Probability that if a person has a weapon, then detector rings
- Probability that if the person is sick, then the test is positive
- Test that always says "yes" has 100% sensitivity

Specificity:
percentage of correct **negatives**

- $\Pr(\bar{A}|\bar{B}) = \Pr(\bar{A} \cap \bar{B}) / \Pr(\bar{B})$
- Probability that if the detector rings, then the person has a weapon
- Probability that if the person is not sick, then the test is negative
- Test that always says "no" has 100% specificity.

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Sensitivity and specificity

Sensitivity: percentage of correct (true) positives
 $\Pr(A|B) = \Pr(A \cap B) / \Pr(B)$

Specificity: percentage of correct (true) negatives
 $\Pr(\bar{A}|\bar{B}) = \Pr(\bar{A} \cap \bar{B}) / \Pr(\bar{B})$

- Let green area be what test detected as positive (red as negative)
 - Event A: in green area. Event B: "+"
- The blue line represents correct separation between positives (+) and negatives (-)
 - Positive can mean "sick", negative "healthy"
- The green part under blue line is false positives.
- The red part above blue line is false negatives.

- 24 + (positives),
- 30 - (negatives)
- 23 true positives
- 22 true negatives
- 8 false positives
- Sensitivity: $100 \cdot 23 / 24 = 96\%$
- Specificity: $100 \cdot 22 / 30 = 73\%$

- One false negative

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Medical test problem



- Consider a medical test that checks for a disease. This test
 - Has false positive rate of 3% (healthy labeled as sick)
 - Specificity** 97%
 - Has false negative rate of 1% (sick labeled as healthy).
 - Sensitivity** 99%
- What is the probability that a person has the disease given that the test came positive?
 - Let A: person tested positive, B: person is sick. $\Pr(B|A)$?
 - $\Pr(A|B) = 0.99$, $\Pr(\bar{A}|B) = 0.01$...
- Not enough information!

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Bayes theorem



- Bayes theorem** allows us to get $\Pr(B|A)$ from $\Pr(A|B)$, if we know probabilities of A and B:

$$\Pr(B|A) = \frac{\Pr(A|B) \cdot \Pr(B)}{\Pr(A)} = \frac{\Pr(A|B) \cdot \Pr(B)}{\Pr(A|B) \cdot \Pr(B) + \Pr(A|\bar{B}) \cdot \Pr(\bar{B})}$$

- Proof:

$$\begin{aligned} \Pr(A|B) &= \frac{\Pr(A \cap B)}{\Pr(B)}, & \Pr(B|A) &= \frac{\Pr(A \cap B)}{\Pr(A)} \\ \Pr(A \cap B) &= \Pr(A|B) \Pr(B) = \Pr(B|A) \Pr(A) \\ \text{So } \Pr(B|A) &= \Pr(A|B) \Pr(B) / \Pr(A) \end{aligned}$$

- The formula $\Pr(A) = \Pr(A|B) \Pr(B) + \Pr(A|\bar{B}) \Pr(\bar{B})$ comes from writing probability of A (e.g., a positive test) as sum of probabilities of A for B (sick people) and for \bar{B} (healthy people).

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$$\Pr(B|A) = \frac{\Pr(A|B) \cdot \Pr(B)}{\Pr(A)}$$

- Consider a medical test that
 - Has false positive rate of 3% (healthy labeled as sick).
 - Has false negative rate of 1% (sick labeled as healthy).
 - Tests for a disease that occurs in 5 in 1000 people.
- What is the probability that a person is sick given that the test came positive?
 - Let A: person tested positive, B: person is sick. $\Pr(B|A)$?
 - $\Pr(A|B) = 0.99$, $\Pr(\bar{A}|B) = 0.01$. $\Pr(\bar{A}|\bar{B}) = 0.97$, $\Pr(A|\bar{B}) = 0.03$
 - $\Pr(B) = 0.005$.
 - $\Pr(A) = \Pr(A|B)\Pr(B) + \Pr(A|\bar{B})\Pr(\bar{B}) = 0.0348$
 - By Bayes theorem, $\Pr(B|A) = \frac{\Pr(A|B) \Pr(B)}{\Pr(A)} = 0.1422$
- So the probability that a person who tested positive has the disease is just 0.1422, that is, 14.22%.
 - By a similar argument, probability that a person who tested negative does not have a disease is whopping $0.99995 = 99.995\%$.

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Hat check problem

- In Victorian England, n men came to an event, and checked in their hats at the door.
 - On the way out, in a hurry, they each picked up a random hat.
 - On average, how many men picked their own hat?



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Expectations

- Often we are interested in what outcome we would see "on average".
 - How fast does this program run "on average"?
- Let possible outcomes of an experiment be numbers a_1, \dots, a_n
 - E.g., time a program takes to sort n elements
- Its **expected value (mean)** is $\sum_{k=1}^n a_k \Pr(a_k)$
 - Often phrased in terms of a "random variable" X, where X is a function from outcomes to numbers.
 - Write $E(X)$ to mean the expected value (mean, expectation) of X.

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Random variables

- “Random variable” X is a *function* from outcomes to numbers.
 - In Computer Science applications, usually X counts something.
 - Number of heads out of n coin tosses.
 - Number of steps a program takes on an input
 - An *indicator* random variable X takes value 0 or 1 depending on whether an event occurred or not.

- Expectation of a random variable X is

$$E(X) = \sum_{i \in \text{Outcomes}} X(i) * \Pr(i)$$



Expectation example

Suppose we roll two fair dice. What is the expected sum of their values?

- X can take values from 2 to 12.
 - $\Pr(X=2) = \Pr(X=12) = 1/36$
 - $\Pr(X=3) = \Pr(X=11) = 2/36 = 1/18$,
 - $\Pr(X=4) = \Pr(X=10) = 3/36 = 1/12$,
 - $\Pr(X=5) = \Pr(X=9) = 4/36 = 1/9$,
 - $\Pr(X=6) = \Pr(X=8) = 5/36$,
 - $\Pr(X=7) = 6/36 = 1/6$



$$E(X) = 2 \cdot \frac{1}{36} + 3 \cdot \frac{1}{18} + 4 \cdot \frac{1}{12} + 5 \cdot \frac{1}{9} + 6 \cdot \frac{5}{36} + 7 \cdot \frac{1}{6} + 8 \cdot \frac{5}{36} + 9 \cdot \frac{1}{9} + 10 \cdot \frac{1}{12} + 11 \cdot \frac{1}{18} + 12 \cdot \frac{1}{36} = 7$$

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Expected win in a lottery

- Rules of Lotto 6/49:
 - A player chooses 6 numbers, 1 to 49.
 - During a draw, 6 randomly generated numbers are revealed.
 - If all 6 numbers chosen by the player match 6 numbers in the draw, the player gets the jackpot of \$5,000,000 or more.
 - There are also smaller prizes; let's ignore them for simplicity.
 - A ticket costs \$3.
 - According to Atlantic Lotto Corporation, chances of winning the jackpot are roughly $1/14,000,000$.

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Expected win in a lottery

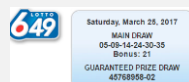
- What is the expected amount a player would win if the jackpot is 5,000,000?
 - $\Pr(\text{win}) = 1/14,000,000$. $\Pr(\text{loss}) = 1 - \Pr(\text{win}) = 13,999,999/14,000,000$.
 - Let the random variable X encode the amount a player wins.
 - For all but one player, that amount is -3. So $\Pr(X=-3) = \Pr(\text{loss})$
 - For the lucky one, the amount is the jackpot minus ticket price. $\Pr(X=4,999,997) = \Pr(\text{win})$
 - Expected amount to win is $E(X) = \Pr(\text{loss}) * (-3) + \Pr(\text{win}) * (5,000,000 - 3) = -2.64$
 - If counting smaller prizes, just add their amount * odds to the sum, and adjust $\Pr(\text{loss})$
 - $E(X) = \Pr(\text{loss}) * (-3) + \Pr(\text{jackpot}) * (4,999,997) + \Pr(5/6 + \text{bonus}) * 374,997 + \Pr(5/6) * 312,497 \dots$

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Expected win in a lottery

- How large should be the jackpot so that the players expect at least to break even?
 - Let's call the jackpot amount J .
 - Expected amount to win is $E(X) = \Pr(\text{loss}) * (-3) + \Pr(\text{win}) * (J - 3)$.
 - To break even, want $E(X) = 0$.
 - $J = 3 + (E(X) - \Pr(\text{loss}) * (-3)) / \Pr(\text{win}) = 42,000,000$




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Bernoulli trials and repeated experiments

- Suppose an experiment has two outcomes, 1 and 0 (success/failure), with $\Pr(1) = p$.
 - Such experiment is called a **Bernoulli trial**.
- What happens if the experiment is repeated multiple times (independently)?



 - A sample space after carrying out n Bernoulli trials is a set of all possible n -tuples of elements in $\{0,1\}$ (or {success, fail}).
 - Number of n -tuples with k 1s is $\binom{n}{k}$
 - Probability of getting 1 in any given trial is p , of getting 0 is $(1-p)$.
 - Probability of getting exactly k 1s (successes) out of n trials is $\binom{n}{k} p^k (1-p)^{n-k}$
 - Probability of getting the first success on exactly the k^{th} trial is $p(1-p)^{k-1}$
- How many trials do we need, on average, to get a success?

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Expected number until...

- Suppose that Alan insists on buying lottery tickets until he wins. If probability of winning is p , how many tickets in expectation ("on average") would he have to buy?
 - Let X be a random variable for how many tickets he has to buy.
 - The probability of winning on exactly i^{th} ticket is $p(1-p)^{i-1}$
 - $E(X) = \sum_{i \in \mathbb{N}} i * \Pr(X = i) = \frac{1}{p}$
 - So for Lotto 6/49 he'd have to buy 14,000,000 tickets (and spend \$42,000,000 – that's jackpot that would let him break even!)

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Expected number until...

- Suppose we have Bernoulli trials with success probability p . What is the expected number of trials to see success?
 - Let X be a random variable for the number of steps till success.
 - $E(X) = \sum_{i \in \mathbb{N}} i * \Pr(X = i) = \frac{1}{p}$
- Same reasoning applies to other processes, where there is a fixed probability of something happening at each experiment or time step.
 - A system has a 1% probability of hanging in any given hour. How long, on average, will it stay up?
 - 100 hours: a little over 4 days.

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Linearity of expectation

- Expectation is a very well-behaved operation:
 - $E(X_1 + X_2 + \dots + X_n) = E(X_1) + \dots + E(X_n)$
 - $E(aX + b) = a E(X) + b$
 - Where $X_1 \dots X_n$ are random variables on some sample space S , and $a, b \in \mathbb{R}$
- Proof:
 - $E(X_1 + X_2) = \sum_{s \in S} p(s)(X_1(s) + X_2(s))$

$$= \sum_{s \in S} p(s)X_1(s) + \sum_{s \in S} p(s)X_2(s) = E(X_1) + E(X_2)$$
 - Similar for $E(aX + b) = a E(X) + b$
 - Using the fact that $\sum_{s \in S} p(s) = 1$

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Hat check problem

- In Victorian England, n men came to an event, and checked in their hats at the door.
 - On the way out, in a hurry, they each picked up a random hat.
 - On average, how many men picked their own hat?



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Hat check problem



In Victorian England, n men came to an event, and checked in their hats at the door.
 – On the way out, in a hurry, they each picked up a random hat.
 – On average, how many men picked their own hat?

- For each man, introduce a random variable X_i , where $X_i = 1$ iff he picked his own hat
 - Such random variables are called **indicator variables**.
 - The quantity we want is $E(X_1 + \dots + X_n)$
 - Now, for each X_i , $E(X_i) = 1 \cdot \Pr(X_i = 1) = \frac{1}{n}$
 - By linearity of expectation, $E(X_1 + \dots + X_n) = E(X_1) + \dots + E(X_n) = \frac{1}{n} + \dots + \frac{1}{n} = n \cdot \frac{1}{n} = 1$
- So on average, just one man will go home with his own hat!

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Hat check problem 2

- In Victorian England, n men came to an event, and checked in their hats at the door.
 - On the way out, in a hurry, they each picked up a random hat.
 - What is the probability that at least k of them got their own hat?



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Deviation from the expectation



“three-quarters (73 percent) of U.S. drivers consider themselves better-than-average drivers” [AAA].

- Let R be a (nonnegative) random variable. What is the probability that $R \geq x$?
- **Markov's inequality:** $\Pr(R \geq x) \leq \frac{E(R)}{x}$
 - Since R is nonnegative, $E(R) \geq x \Pr(R \geq x) + 0 \Pr(R < x)$
 - So $E(R) \geq x \Pr(R \geq x)$.

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Using Markov's inequality



- **Markov's inequality:** $\Pr(R \geq x) \leq \frac{E(R)}{x}$
- Let's use Markov's inequality in the Hat Check Problem
 - Let $R = X_1 + \dots + X_n$
 - each X_i is an indicator variable for i^{th} man getting his hat.
 - By Markov's inequality, $\Pr(R \geq k) \leq \frac{E(R)}{k} = \frac{1}{k}$.
- But that's a pretty weak bound!
 - Probability that everybody gets their hat is $\Pr(R \geq n) \leq \frac{1}{n}$ by Markov's inequality
 - However, we can calculate that it is only $\frac{1}{n!}$

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Variance



Let's define the expectation of the difference from the mean.

Let X be a random variable on a sample space S . Define **variance** of X , denoted $V(X)$ to be

$$V(X) = E_{s \in S} \left((X - E(X))^2 \right) = \sum_{s \in S} (X(s) - E(X))^2 \cdot \Pr(s)$$

A **standard deviation** of X , denoted $\sigma(X)$ where σ is a Greek letter sigma, is $\sigma(X) = \sqrt{V(X)}$

Useful equality: $V(X) = E(X^2) - (E(X))^2$

- That is, the variance of X can be expressed as the difference between the expectation of X^2 and square of $E(X)$.
- To prove this, just open parentheses in our original definition of $V(X)$, and use the fact that $\sum_{s \in S} \Pr(s) = 1$.

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Variance



Variance of a random variable X on a sample space S is:

$$V(X) = E_{\text{res}}((X - E(X))^2) = \sum_{s \in S} (X(s) - E(X))^2 \cdot \Pr(s) = E(X^2) - (E(X))^2$$

A standard deviation of X is $\sigma(X) = \sqrt{V(X)}$

Example: consider again throwing a fair die

- and let X be a random variable for the number that comes up
 - So X takes values $1, 2, 3, 4, 5, 6$ with equal probability $1/6$.

$$E(X) = \sum_{r \in \{1, 2, 3, 4, 5, 6\}} r \cdot \frac{1}{6} = \frac{7}{2}$$

$$E(X^2) = \sum_{r \in \{1, 2, 3, 4, 5, 6\}} r^2 \cdot \frac{1}{6} = \frac{1}{6}(1^2 + 2^2 + \dots + 6^2) = \frac{91}{6}$$

$$V(X) = \frac{91}{6} - \left(\frac{7}{2}\right)^2 = \frac{35}{12} \quad \text{And } \sigma(X) = \sqrt{\frac{35}{12}} \approx 1.7$$



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Variance in the hat check problem



Variance of a random variable X on a sample space S is: $V(X) = E(X^2) - (E(X))^2$

- Example: variance in the hat check problem.

- Let $R = X_1 + \dots + X_n$ for indicator variables X_i : man i gets his hat.
- $E(R) = 1$.

$$E(R^2) = E((\sum_{i=1}^n X_i)^2) = E(\sum_{i=1}^n X_i^2 + \sum_{i \neq j} \sum_{j=1}^n X_i X_j) = E(\sum_{i=1}^n X_i^2) + E(\sum_{i \neq j} \sum_{j=1}^n X_i X_j)$$

$$E(X_i^2) = 1^2 \cdot \frac{1}{n} = \frac{1}{n}, \text{ so } E(\sum_{i=1}^n X_i^2) = \sum_{i=1}^n E(X_i^2) = 1$$

$$E(X_i X_j) = \frac{1}{n(n-1)}, \text{ so } E(\sum_{i \neq j} \sum_{j=1}^n X_i X_j) = (n^2 - n) \cdot \frac{1}{n(n-1)} = 1$$

$$\text{So } E(R^2) = E((\sum_{i=1}^n X_i)^2) = 1 + 1 = 2.$$

$$\text{Finally, } V(R) = E(R^2) - (E(R))^2 = 2 - (1)^2 = 1.$$

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Independent random variables

- Similar to independent events:

- Random variables X, Y are **independent** iff

$$\forall r_1, r_2 \in \mathbb{R} \quad \Pr(X = r_1 \wedge Y = r_2) = \Pr(X = r_1) \cdot \Pr(Y = r_2)$$



- Example: Suppose a pair of fair dice is thrown, with

X_1 : number on the first die, X_2 : number on the second die, Y : total amount

- Then $\Pr(X_1 = i \wedge X_2 = j) = \Pr(X_1 = i) \cdot \Pr(X_2 = j) = \frac{1}{36}$ for all $i, j \in \{1, 2, 3, 4, 5, 6\}$ which are the only values X_1 and X_2 take. So X_1 and X_2 are independent.

- But X_1 and Y are not independent: $\Pr(X_1 = 1 \wedge Y = 12) = 0$, whereas $\Pr(X_1 = 1) = \frac{1}{6}$, and $\Pr(Y = 12) = \frac{1}{36}$, so $\Pr(X_1 = 1) \cdot \Pr(Y = 12) = \frac{1}{216} \neq 0$.

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Variance of Bernoulli trials

- When random variables X, Y are independent, then

$$E(XY) = E(X) \cdot E(Y)$$

$$V(X + Y) = V(X) + V(Y)$$



- With this, we can calculate the variance of the number of successes in n Bernoulli trials where in each trial the probability of success is p .

- For a single trial: let $X = 1$ iff success, $X = 0$ iff failure.

- Then $E(X) = p$, $E(X^2) = p$, so $V(X) = E(X^2) - (E(X))^2 = p - p^2 = p(1 - p)$

- Let $X_1 \dots X_n$ be indicator variables so that $X_i = 1$ iff i^{th} trial is a success.

- They are independent, so $V(X_1 + \dots + X_n) = \sum_{i=1}^n V(X_i) = np(1 - p)$

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Chebyshev's inequality



- Now that we have the notion of variance, we can get a better bound on the probability of being away from the mean:

- Chebyshev's inequality:** Let $x > 0$. Then

$$\Pr(|R - E(R)| \geq x) \leq \frac{V(R)}{x^2}$$

Proof:

- $|R - E(R)| \geq x$ is equivalent to $(R - E(R))^2 \geq x^2$.

- By Markov's inequality, $\Pr((R - E(R))^2 \geq x^2) \leq \frac{E((R - E(R))^2)}{x^2} = \frac{V(R)}{x^2}$

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Better estimate for Hat Check



- Chebyshev's inequality:** Let $x > 0$. Then

$$\Pr(|R - E(R)| \geq x) \leq \frac{V(R)}{x^2}$$

- Example: hat check problem again.

- Let $R = X_1 + \dots + X_n$ for indicator variables X_i : man i gets his hat.

- We have shown that $E(R) = 1$, and that $V(R) = 1$ as well.

- Let $k \geq 2$. Since $R \geq 0$, $\Pr(R \geq k) = \Pr(|R - 1| \geq k - 1) \leq 1/(k - 1)^2$

- This gives a tighter bound than Markov's inequality.

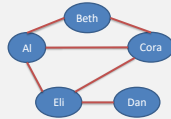
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Team selection puzzle



- Imagine that your friend is a project manager, and her team consist of great programmers – if only she could get them to stop fighting among each other!
 - She decides to split them in two smaller teams
 - To minimize fighting within each team.
 - She knows who fights with whom (the “CONFLICT relation”), but how can she do the splitting?
 - And is it possible at all to eliminate at least half the conflicts? If not, why bother...
- Do you think it is possible to split any group into two teams
 - to eliminate all conflicts?
 - How about eliminating half the conflicts?
 - How would you do the splitting?

- Suppose this is the graph of the CONFLICT relation for a group.
 - Here, lines are double-direction arrows, since CONFLICT is symmetric.



- What do you think is the best split?